Some basic statistics and curve fitting techniques

Some important concepts:

- Data
- Statistical description of data (data reduction, independence)
	- The use of statistics to make a point:
		- 1. Statistics never proves a point.
		- 2. If you need fancy statistic to support a point, your point is, at best, weak…

Statistical description of data Statistical moments (1st and 2nd):

- Mean: $\overline{x} = \frac{1}{N} \sum_{j=1}^{N}$ *N j* x_j *N x* 1 1 • variance: $Var = \frac{1}{N} \sum_{i} (x_i - \overline{x})$ 2 $=\frac{1}{N-1}\sum_{j=1}^{N}(x_j -$ *N j* $x_j - \overline{x}$ *N Var*
- Standard deviation: $\sigma = \sqrt{Var}$
- Average deviation:

$$
Adev = \frac{1}{N} \sum_{j=1}^{N} \left| x_j - \overline{x} \right|
$$

• Standard error:

 $S_{error} = \sigma / \sqrt{N}$

• Standard error:

When is the uncertainty not reduced by sampling more?

Statistical description of data

Probability distribution:

Fig. 1-2 Histogram of frequency distribution of stature of 24,404 U.S. Army males. Adapted from data of Newman and White.

Non-normal probability distribution:

Fig. 1-3 U.S., female, 1965: percent dying in each 5-year age interval (the 100-105 interval includes all deaths after 100 rather than only those occurring in the interval). Data from N. Keyfitz and W. Flieger, World Population: An Analysis of Vital Data. Chicago: University of Chicago Press, 1968, p. 45.

Statistical description of data

Nonparametric statistics (when the distribution is unknown):

• rank statistics

- •Median
- percentile
- Deviation estimate
- The mode

Issue: *robustness*

Statistical description of data

Robust: "insensitive to small departures form the idealized assumptions for which the estimator is optimized."

Figure 14.6.1. Examples where robust statistical methods are desirable: (a) A one-dimensional distribution with a tail of outliers; statistical fluctuations in these outliers can prevent accurate determination of the position of the central peak. (b) A distribution in two dimensions fitted to a straight line; non-robust techniques such as least-squares fitting can have undesired sensitivity to outlying points.

Statistical description of data Examples from COBOP:

Relationship between 2 variables Linear correlation:

$$
r = \frac{\sum_{i} (x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum_{i} (x_i - \overline{x})^2} \sqrt{\sum_{i} (y_i - \overline{y})^2}}
$$

Rank-order correlation:

$$
r_s = \frac{\sum_i (R_i - \overline{R})(S_i - \overline{S})}{\sqrt{\sum_i (R_i - \overline{R})^2} \sqrt{\sum_i (S_i - \overline{S})^2}}
$$

Regressions of type I and type II Uncertainties in y only:

$$
y(x) = ax + b
$$

$$
\chi^2 = \sum_{i=1:N} \left(\frac{y_i - a - bx_i}{\sigma_i} \right)^2
$$

Minimize χ^2 by taking the derivative of χ^2 wrt *a* and *b* and equal it to zero.

What if we have errors in both x and y?

Press et al.'s approach:

$$
y(x) = ax + b
$$

\n
$$
\chi^{2} = \sum_{i=1:N} \frac{(y_{i} - ax_{i} - b)^{2}}{\sigma^{2} y_{i} + a^{2} \sigma^{2} x_{i}}
$$

\n
$$
Var(y_{i} - ax_{i} - b) = \sigma^{2} y_{i} + a^{2} \sigma^{2} x_{i}
$$

Minimize χ^2 by taking the derivative of χ^2 wrt *a* and *b* and equal it to zero.

The coefficient of determination

$R^2 = 1 - MSE/Var(y)$.

MSE=mean square error=average error of model^2/variance.

What variance does it explain?

Can it reveal cause and effect?

How is it affected by dynamic range?

R is the 'correlation coefficient'.

Regressions of type I and type II Classic type II approach (Ricker, 1973):

The slope of the type II regression is the geometeric mean of the slope of y vs. x and the inverse of the slope of x vs. y.

 $y(x) = ax + b$ $\pm = sign\{\sum_{i} x_i y_i\}$ $a_{II} = \sqrt{a/c} = \pm \sigma_{y}/\sigma_{x}$ $x(y) = cy + d$ $sign\{\sum x_i y\}$

Smoothing of data

- Filtering noisy signals.
- What is noise?
- instrumental (electronic) noise.
- Environmental 'noise'.
	- " one person ' s *noise* may be another person ' s *signal*"
	- Matlab: filtfilt

Method of fluctuation

Modeling of data

Condense/summarize data by fitting it to a model that depends on adjustable parameters.

Example, CDM spectra:

$$
a_g(\lambda) = \tilde{a}_g \exp(-s(\lambda - \lambda_0))
$$

particulate attenuation spectra:

$$
c_p(\lambda) = \widetilde{c}_p\left(\frac{\lambda}{\lambda_0}\right)^{-\gamma}
$$

Modeling of data Example: CDM spectra.

$$
a_g(\lambda) = \widetilde{a}_g \exp(-s(\lambda - \lambda_0))
$$

\n
$$
\Rightarrow a = [\widetilde{a}_g, s]
$$

Merit function:

$$
\chi^{2} = \sum_{i=1}^{9} \left[\frac{a_{g}(\lambda_{i}) - \widetilde{a}_{g} \exp(-s(\lambda - \lambda_{0}))}{\sigma_{i}} \right]^{2}
$$

•For non-linear models, there is no guarantee to have a single minimum. •Need to provide an initial guess.

Matlab: fminsearch

Modeling of data Lets assume that we have a model

 $y = y(\lambda; a)$

A more robust merit function:

 $\sum_{i=1}^{N} \left| \frac{y(\lambda_i) - y(\lambda_i; a)}{\sigma_i} \right|$ $\overline{=}$ σ λ_i) – $y(\lambda)$ $\widetilde{\chi}$ = *N* $\overline{i=1}$ $\qquad \qquad$ \qquad \qquad $y(\lambda_i) - y(\lambda_i)$ 1 $\approx -\frac{N}{N}y(\lambda_i)-y(\lambda_i; a)$

Problem: derivative is not continuous. Can be used to fit lines.

Statistical description of data

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Monte-Carlo/Bootstrap methods

Need to establish confidence intervals in:

- 1. Fitting-model parameters (e.g. CDM fit).
- 2. Model output (e.g. Hydrolight).

When there is an uncertainty (or possible error) associated with the input:

Vary inputs with random errors and observe effect on output:

Example: how to assign uncertainties in derived spectral slope of CDOM.

Merit function:

$$
\chi^{2} = \sum_{i=1}^{9} \left(a_{g} \left(\lambda_{i} \right) \pm \Delta_{i} - \tilde{a}_{g} \exp \left(-s \left(\lambda - \lambda_{0} \right) \right) \right)^{2}
$$

Randomly add uncertainties (Δ_i) to each measurement, each time performing the fit (e.g. using randn.m in Matlab, RAND in Excel).

Then do the stats for the different *s.*